4 MODEL BUILDING

When the data was analysed the group which showed a significant difference in tumour response rates was the comparison of Japanese and non-Japanese patients. To explore the reason for this apparent difference the data was analysed using logistic regression. The first analysis did not account for any baseline factors other than ethnicity and this resulted in an odds ratio of 3.27, indicating that the chances of responding was over 3 times higher for Japanese patients compared with non-Japanese patients (Table 1).

Table 1 Unadjusted Model

Parameter	Odds 95% CI Ratio	p-value	Interpretation
Ethnicity	3.27 1.57, 7.26	0.0023	The odds of responding is over 3 times higher for Japanese patients compared to non-Japanese patients.

CI Confidence interval.

In order to account for the observed baseline imbalances seen between Japanese and non-Japanese patients further logistic modelling was performed. This allowed odds ratios to be calculated from the model parameters, but unlike simple 2 x 2 tables the odds ratios were adjusted for all other relevant factors in the model. Therefore, the methodology allows the variation in the data to be explored further, making the assessment of the ethnic difference more sensitive and accurate.

Before the modelling was performed the data was reviewed to identify clinically meaningful baseline factors that may influence tumour response. The factors were then made into binary factors (0 or 1) or continuous factors. Each of the factors were then analysed in isolation to assess whether they were predictive of response. Those factors found to be of predictive of response at the 0.10 level were then considered in the multivariate logistical analysis. Table 2 shows the p-value for each of the parameters tested in the modelling.

Table 2 Model Building – univariate effects

	p-value
Parameter	0.9553
Duration of previous chemotherapy treatn	nemi
Months from diagnosis to randomisation	
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Age group (<65 years vs ≥65 years)	0.7005

ALCONOMIC

Parameter	p-value
Type of disease (measurable/non-measurable)	0.5280
Stage of disease (III vs IV)	0.4530
Number of evaluable lesions at entry	0.4342
Number of measurable lesions at entry	0.4325
Progressed on a previous chemotherapy	0.3522
Time from last dose of chemotherapy to randomisation	0.3156
Visceral metastases at entry	0.1838
Previously received surgery	0.1658
Tumour burden at entry	0.1512
History of lung disorder, chest pain, dyspnoea, increased cough or	0.1413
haemoptysis: more control beauty majorating incoming	
Previously received docetaxel	0.1103
Baseline lung cancer subscale score	0.0923ª
Body mass index at entry	0.0887ª
Performance status	0.0619ª
Previously received radiotherapy	0.0587ª
Histology wilds statement in the program of the principal program of the control	0.0013 ^a
Previously received other treatment ^b	0.0004
Gender	0.0003ª

^a p<0.10: significance level for inclusion in the model (as stated in protocol).

As shown in Table 2, the baseline factors found to be predictive of response in isolation were baseline lung cancer subscale score, BMI, PS, receipt of previous radiotherapy, tumour histology, gender, and receipt of previous other treatment. Although the significance level used for model building was 0.1, as stated in the protocol, a further analysis was done using a 0.15 level to assess the robustness of the model. Using the higher threshold, two more factors were included in the logistic model (see Table 2). However, when the factors were considered in further multivariate models they were rejected at the 0.15 significance level, thus resulting in the same final model as found using a 0.1 threshold level.

The next step was to fit these seven parameters in one logistical model to assess their impact on the apparent difference seen between the ethnic groups. By incorporating this information into

b Other treatments include picibanil, investigational drugs, minomycin, marimastat and NOLVADEX.

one model, it allowed the ethnic comparison to be assessed after controlling for prognostic factors (see Table 3).

Table 3 Model Building - multivariate effects

Parameter				p-value
Body mass index at ent	ıy			0.7889
Previously received rac				0.6766
Ethnicity		i		0.2530
Baseline lung cancer su	ibscale score		1664	28420,2231 科理
Performance status	ngan ganamanan sasar sa Marini sangangan	es es es es estados. Maistras estados estad		0.0814 ^a
Histology				0.0212ª
	cade fee vide out	4,747.5	No. of the April 1	0.0166ª
Previously received of	her treatment ^a			0.0108 ^a

^a p<0.10: significance level for inclusion in the model (as stated in protocol).

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5 FINAL MODEL

As shown in Table 3, the main effects model indicated that PS, histology, gender and receipt of other treatments were related to tumour response. Although ethnicity was not significant at the 10% level, it was retained in the model to allow a final assessment of ethnic difference after adjustment for prognostic factors. The final step in the modelling was to assess whether there were any interactions between the prognostic factors. However, no interactions were significant (p>0.4), so the main effects model was considered to be the best interpretation of the data (Table 4).

Table 4 Final Adjusted Model

Parameter	Odds Ratio	95% CI	p-value	Interpretation https://doi.org/10.1001
Performance status	6.26	1.20, 115.36	0.0814	The odds of responding is over 6 times higher for PS 0 or 1 patients compared to PS 2 patients.
Received prior other treatment ^a	6.01	1.58, 26.15	0.0108	The odds of responding is 6 times higher for patients who received other treatments* prior to entry compared to those who did not.
Histology	3.45	1.29, 11.02	0.0212	The odds of responding is almost 3 ½ times higher for patients with adenocarcinoma compared to patients with other turnour histologies.
Gender	2.65	1.19, 5.91	0.0166	The odds of responding is over 2 $\frac{1}{2}$ times higher for females than males.
Ethnicity	1.64	0.71, 3.93	0.2530	After accounting for all baseline imbalances the odds ratio indicates that the chance of responding is just over 1½ times higher for Japanese patients compared to non-Japanese patients.

^a Other treatments include picibanil, investigational drugs, minomycin, marimastat and NOLVADEX. CI Confidence interval.

The final column of Table 4 provides an explanation of the results. By comparing the model without adjustment for prognostic factors to the model with adjustment for prognostic factors, it was clear the amount of variation explained by these variables. Without the variation being explained in the unadjusted model (Table 1), the odds ratio for ethnicity was 3.27 (p=0.0023).

PS Performance status.

However, after including these variables in the model, and allowing a more accurate assessment of the ethnic difference, the odds ratio was halved to 1.64 (p=0.2530).

From the modelling results, it can be concluded that the odds of responding is 1.64 times higher for Japanese patients compared to non-Japanese patients, but as the 95% confidence interval crosses the value of 1 (representing equality) this difference is not considered to be statistically significant (p=0.2530).

Using the following logit model and the parameterisation shown in Table 5, it was possible to calculate estimated probabilities of response for individual patients. This was done by substituting the relevant value of x_k (ie, either 0 or 1) into the equation below:

logit (p) =
$$-4.8978 + 0.4951 * x_{\text{ethnicity}} + 1.8341 * x_{PS} + 1.7930 * x_{\text{other}} + 0.9726 * x_{\text{gender}} + 1.2382 * x_{\text{histology}}$$

Table 5 Parameterisation for logistic model

Parameter	Flags
$\mathcal{X}_{ ext{edimicity}}$	0≔non-Japanese 1=Japanese
x_{PS}	0=PS 2 1=PS 0 or 1
x_{other}	0=did not receive other previous treatment 1=did receive previous other treatment
$x_{ m gender}$	0=male 1=female
$x_{ m histology}$	0=squamous, undifferentiated, large cell or squamous & adenocarcinoma 1=adenocarcinoma

PS Performance status.

If we were to use the model to compare the probability of response for a Japanese patient given the average baseline characteristics of a non-Japanese patient (ie, PS=0-1, no other treatments, male and having adenocarcinoma), then we would find that the predicted probability of response was 20.9%. In a similar fashion, if we were to use the model to compare the probability of response for a non-Japanese patient given the average baseline characteristics of a Japanese patient (ie, PS=0-1, no other treatments, male and having adenocarcinoma), then we would find that the predicted probability of response was 13.9%.

a Paratina.

In addition to this example, the model shows that at the most extreme situations, the estimated probability of response ranged from 0.74% to 71.9% for non-Japanese patients, and 1.21% to 80.8% for Japanese patients. Thus, when all prognostic factors are considered in the modelling, the range of response rates are very similar between the two ethnic groups.

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6 DISCUSSION

Without making any adjustment for baseline imbalances, the odds of responding was over 3 times higher for Japanese patients compared to non-Japanese patients (p=0.0023). However, upon reviewing the data, it was evident that there were many prognostic factors that favoured the Japanese patients. In order to account for these baseline imbalances, logistic modelling was performed to allow a more accurate assessment of the ethnic difference.

After accounting for baseline imbalances, the odds ratio for ethnicity was 1.64 (p=0.2530) suggesting that the chances of responding was 1.64 times higher for the Japanese patients compared with the non-Japanese patients. However, as the confidence interval ranged from 0.71 to 3.93, we could not rule out the possibility that the true odds ratio may be equal to unity, indicating equal response rates in the ethnic groups.

Using the final logistic model, it was possible to calculate the estimated probabilities of response for individual patients depending on whether or not they had the prognostic factors identified in the modelling (ie, PS=0 to 1, receipt of prior other treatment, female, and adenocarcinoma histology). Estimation of the probability of response for a Japanese patient with the average baseline characteristics of a non-Japanese patient, gave a probability of response of 20.9%. Using the same methodology, the probability of response for a non-Japanese patient with the average baseline characteristics of a Japanese patient, gave a probability of response of 13.9%.

These estimated probabilities or response highlight the wide range of results that can be seen between patients irrespective of whether they are Japanese or non-Japanese. However, the fact that this trial involved a large number of patients (n=210), it is unlikely that the results could be heavily influenced by patients with a very poor prognosis or patients with a very good prognosis. The trial data showed that the trial had a large representative population, thus making it likely that the trial results can be reproduced.

7 CONCLUSION

The results have suggested that without adjustment for baseline imbalances between Japanese and non-Japanese groups, there was a large difference between the two ethnicities. However, after accounting for the prognostic factors identified in the trial (ie, PS, histology, gender and the receipt of previous treatments other than chemotherapy, radiotherapy and surgery), using the modelling approach, it was clearly demonstrated that there was no statistically significant difference between the ethnic groups. In addition, when probabilities of response for patients within each ethnic group were estimated, the range of results were hugely overlapping; confirming similarity. This highlighted that when all prognostic factors were considered in the modelling, the range of response rates were similar between the two ethnic groups.

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APPENDIX A

Summary tables produced in response to DO questions

	Response rates and durations of first-line chemotherapy regimen presented by dose
Tables T99.4 to T99.6	Response rates and durations of first-line chemotherapy presented by dose and ethnicity
Tables T99.7 to T99.9	Response rates and durations of second-line chemotherapy presented by dose and ethnicity

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CAUSE OF DEATH
POPULATION: EFS PATIENTS WHO WERE INTOLERANT TO LAST CHEMO REGIMEN & WHO DIED WITHIN 4 MONTHS OF RANDOMISATION

RANDOMISED TREATMENT = GEFITINIE

DEATH	RELATED TO CANCER	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
i.	AUTOPSY DONE	No	No	No	No	No.	No	No	No No	No	No	No	No	Q	Š.	ON ON	N N O
148 m	SECONDARY CAUSE PREFERRED TERM	1 g 3.		ng in I	NON-SMALL CELL	NON-SMALL CELL	LUNG NEOPLASM	NON-SMALL CELL	PNEUMONIA			Bargay Albar Abar	X 4: (4:3) (4:3)		evelle Augus Avena	NON-SMALL CELL LUNG CANCER METASTATIC	HAEMOPTYSIS
	SECONDARY CAUSE OF DEATH				Non-small call	Tung cancer Progression of	Caused by progressive lung		Pneumonia					-		Pulmonary metastases of non small cell lung	cancer Hemoptysis 3
	PRIMARY CAUSE PREFERRED TERM	NON-SMALL CELL LUNG	NON-SMALL CELL LUNG	CANCER NON-SMALL CELL LUNG CANCER	PULMONARY EMBOLISM	RESPIRATORY FAILURE	CARDIOPULMONARY FAILURE	RESPIRATORY FAILURE	MULTI-ORGAN FALLURE SEPSIS	NON-SMALL CELL LUNG	NON-SMALL CELL LUNG	CANCER NON-SMALL CELL LUNG CANCER	NON-SMALL CELL LUNG CANTER	NON-SMALL CELL LUNG	NON-SMALL CELL LUNG CANCER	LUNG NEOPLASM MALICHANT RESPIRATORY FAILURE	RESPIRATORY FAILURE ACUTE RESPIRATORY FAILURE
	TIME TO PRIMARY CAUSE DEATH OF DEATH	1.87 Non small cell lung	1.28 Non-small cell lung	2.53 Non small cell lung	1.25 Pulmonary embolism	0.92 Respiratory insufficiency RESPIRATORY FALLURE	3.25 Kardio - resp insuff	3.29 Respiratory failure	0.79 Multiple organ failure 2.63 Respiratory insuficiency	0.66 Non small cell lung	1.15 Non-small cell lung	1.08 Non small cell lung	1.45 Words small cell lung	3.32 NSCLC	1.12 Non small cell lung cancer - progressive	0.69 Lung cancer progression 1.08 Respiratory insufficiency	1.41 Respiratory insufficiency RESPIRATORY FAILURE 1.94 Acute respiratory insufficiency
	PATIENT	E0113004	E0147002	E0150005	E0341002	E0505018	E0505056	E0505058	E0568004 E0587004	E0622011.	E1108005	E1125008	E1126005	E1165001	E1356004	E1460006	E1461032 E1461056

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CAUSE OF DEATH
POPULATION: EFS PATIENTS WHO WERE INTOLERANT TO LAST CHEMO REGIMEN & WHO DIED WITHIN 4 MONTHS OF RANDOMISATION

RANDOMISED TREATMENT = GEFITINIB

	DEATH RELATED	TO CANCER	တ္	ស្ត	_	Ŋ	ន្ត	Yes	SS	Yes	Yes	Yes	;	· į					ON	Yes	Yes	!	Yes		Yes		Yes
	DE AUTOPSY RE		No Yes	No Yes		10	No Yes	No Ye	No Yes		No Ye	No Ye							Σ Q	No Y	No X		No.	- 1	No		No Y
	A TELESCONDARY CATISE A	TERM	LUNG NEOPLASM NAT. TCNANT	LASM		LUNG NEOPLASM Y MALIGNANT	CARDIOPULMONARY		4	4	A	NON-SMALL CELL	LUNG CANCER STAGE	ΛΤ					13	TORY	FAILURE				***	LUNG CANCER	
= CERTIFIED		SECONDARY CAUSE OF DEATH	Lung cancer	Lung cancer		Lung carcinoma	Cardiorespiratoric	ישריהי				er Bronchonenic/non		cancer stage iv	brain metastases and pleural	effusion (right)	s/p closed tube	<pre>removal (right)</pre>	¥, *	Respiratory	failure				Mon-temp-	lung cancer	
RANDOMISED TREATMENT =		PRIMARY CAUSE PREFERRED TERM	RESPIRATORY FAILURE	MULTI-ORGAN FAILURE		RESPIRATORY FALLONG METASTATIC NEOPLASM	NON-SMALL CELL LUNG	CANCER NON-SMALL CELL LUNG	CANCER NON-SMALL CELL LUNG	TIME CANCER METASTATIC	NON-SMALL CELL LUNG	CANCER	CARDIO-RESPIRATORI PARE				1. 日本教の日報の 古典の書名		DEATH	NON-SMALL CELL LUNG	CANCER	CANCER METASTATIC	Chita transcription	CANCER METASTATIC	DELLITAGE SCHOOLS	RESERVATIONE FALLONE	NON-SMALL CELL LUNG CANCER
		TIME TO PRIMARY CAUSE DEATH OF DEATH	0.43 Respiratory insufficiency RESPIRATORY FAILURE	0,72 Multiple organs collapse	•	 1.38 Respiratory insufficiency 3.19 Carcinomatosis 	3.29 Non small cell lung	cancer 1.74 Progression of subject's	nsclc 3.42 NSCLC progression		3.02 Metastalc lung cancer 3.58 NSCLC	(1) (1) (1) (1) (1) (1) (1) (1) (1) (1)	1.58 Cardiopulmonary arrest	disseminated malignancy.					1.5	expired in a remote piace 2.92 Progression of non small	cell lung cancer	3.29 Metastatic, progressive non-small cell lung	cancer.	1.22 Progressive metastatic non small cell lung	cancer	0.85 Kespiratory raile	1.18 Non small cell lung cancer
		TIME				E1461080 1 E1461087 3					E1733004 3		ES300003 1			3			ES706006	F5804020		E6003008		E6003039		E6108006	E6600001

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CAUGE OF DEATH & WHO WERE INTOLERANT TO LAST CHEMO REGIMEN & WHO DIED WITHIN 4 MONTHS OF RANDOMISATION

	DEATH	TO CANCER	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	÷.
i	1	AUTOPSI	No	No	No No	No	No	No	No	Ño	No	į.
		SECONDARY CAUSE PREFERRED TERM	NON-SMALL CELL					NON-SMALL CELL LUNG CANCER	LUNG NEOPLASM	MALLGNANT	LUNG NEOPLASM	
PLACEBO		SECONDARY CAUSE OF DEATH	Progression of	BSCIC			一年 第二人 医神经神经	Progression of non-small cell	Lung cancer		Lung cancer	
RANDOMISED TREATMENT = PLACEBO	The second secon	PRIMARY CAUSE PREFERRED TERM	RESPIRATORY FAILURE	NON-SMALL CELL LUNG	CANCER NON-SMALL CELL LUNG	CANCER	LUNG NEOPLASM MALIGNANT	LUNG NEUPLASSE EMPLICATION SUPERIOR VENA CAVAL OCCIUSION	RESPIRATORY FAILURE		BRONCHOPNEUMONIA CHRONIC OBSTRUCTIVE	AIRWAYS DISEASE
		TIME TO PRIMARY CAUSE DEATH OF DEATH	0.46 Respiratory failure	2.46 NSCLC	1 September 1 of 100 September 1	cell lung cancer	2.30 Lung cancer	2.99 Lung cancer 1.61 Superior vena cava	2 AE Bilmonawe inenfficiency	The state of the s	0.36 Bronchopneumonia	pulmonary disease
		PATTENT	E0505005	51000015		TOOTETTS	E1173001	E1201001 E1210001		ELABIUS.	E1462003	200